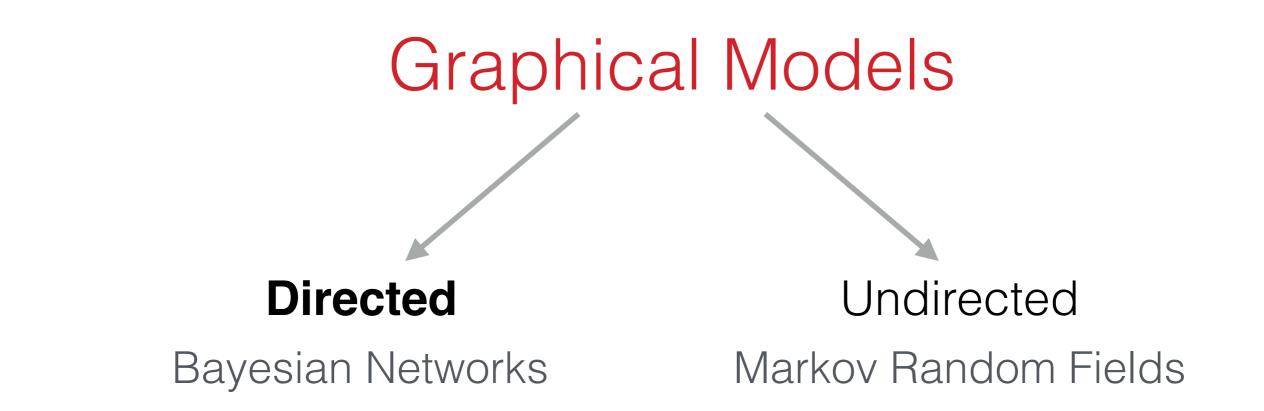
Graphical Models Review

Anthony Platanios



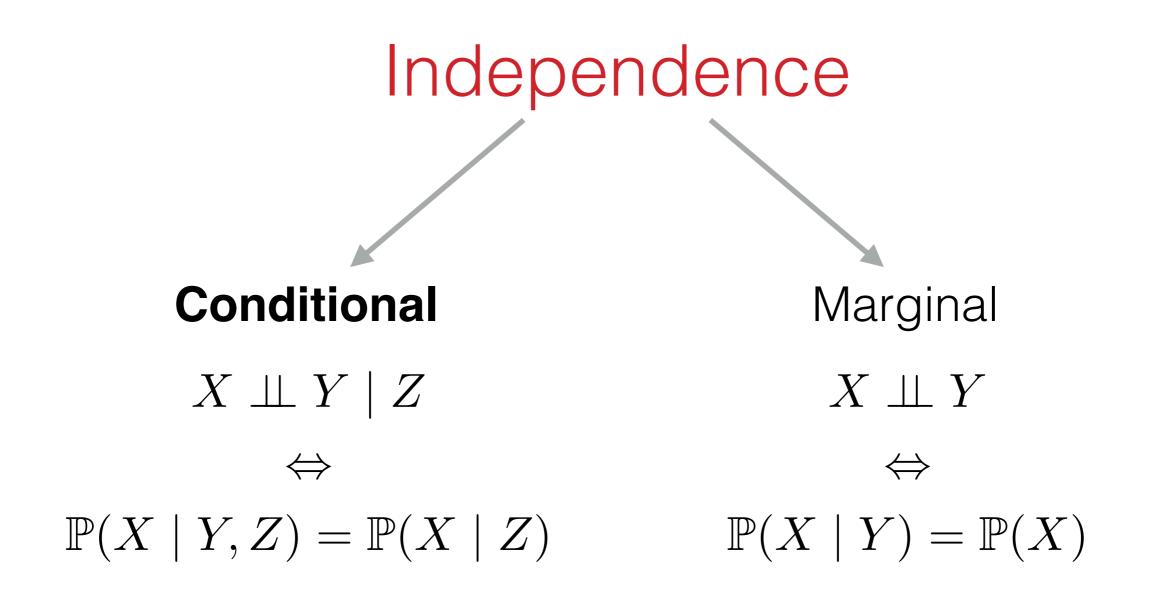
Model

Conditional independence assumptions

• Joint probability distribution of variables (parameterized)

Combine Prior Knowledge

- Over dependencies
- Over parameter values

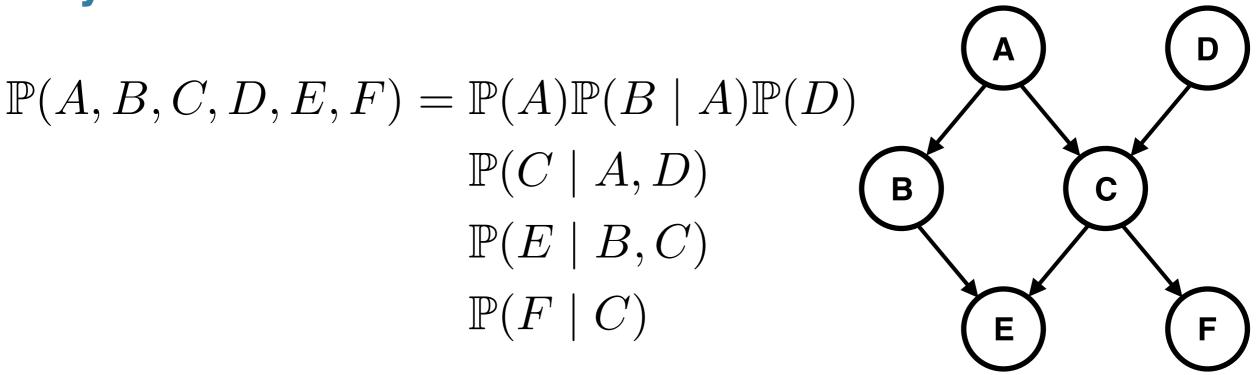


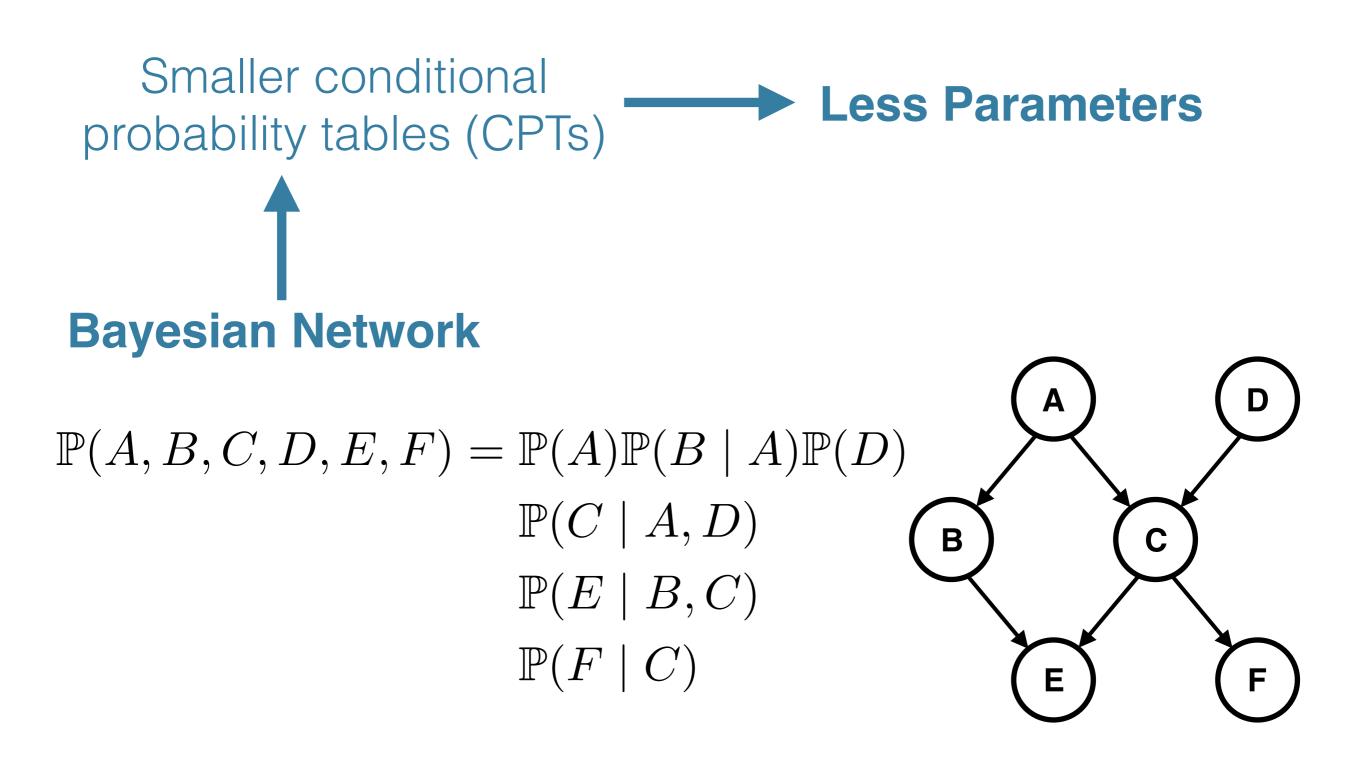
How do directed graphical models help?

Chain Rule of Probability

$$\mathbb{P}(A, B, C, D, E, F) = \mathbb{P}(A)\mathbb{P}(B \mid A)\mathbb{P}(C \mid A, B)$$
$$\mathbb{P}(D \mid A, B, C)\mathbb{P}(E \mid A, B, C, D)$$
$$\mathbb{P}(F \mid A, B, C, D, E)$$

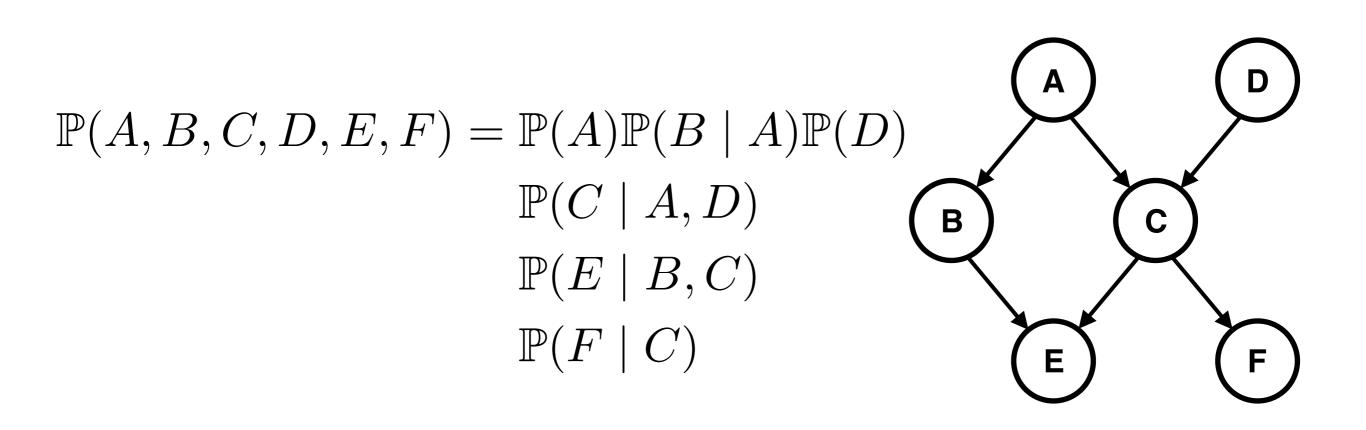
Bayesian Network





Nodes encode variables

Edges encode dependencies



In general:

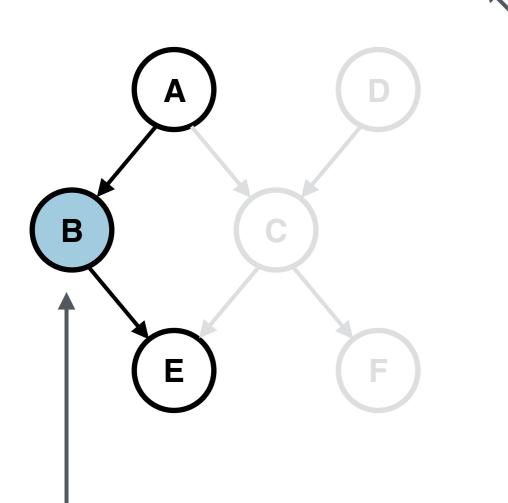
$$\mathbb{P}(X_1, \dots, X_n) = \prod_{i=1}^n \mathbb{P}(X_i \mid \text{Parents}(X_i))$$

But can we determine general conditional independence properties?

Yes, with the **d-separation criterion!**

In order to see how this is possible, let us first consider **three simple cases**. Then we are going to see a simple way (more like a game) for figuring out independence properties of a graph using this criterion.

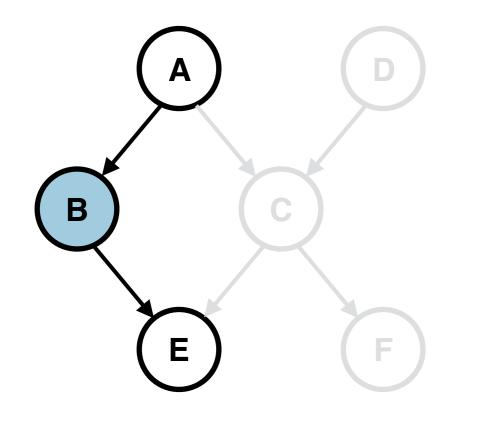
Case #1: Head to Tail



"Heads" and "tails" refer to the connecting edges heads (i.e., arrow pointers) and tails

Shaded nodes are observed and we want to see if observing them induces any independencies

Case #1: Head to Tail



$$\mathbb{P}(A, B, E) = \mathbb{P}(A)\mathbb{P}(B \mid A)\mathbb{P}(E \mid B) \Rightarrow$$

$$\mathbb{P}(A, E \mid B) = \frac{\mathbb{P}(A)\mathbb{P}(B \mid A)\mathbb{P}(E \mid B)}{\mathbb{P}(B)}$$

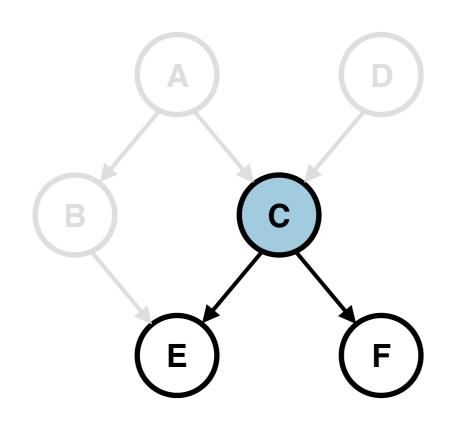
$$= \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}\mathbb{P}(E \mid A)$$

$$= \mathbb{P}(A \mid B)\mathbb{P}(E \mid B)$$

$$\clubsuit$$

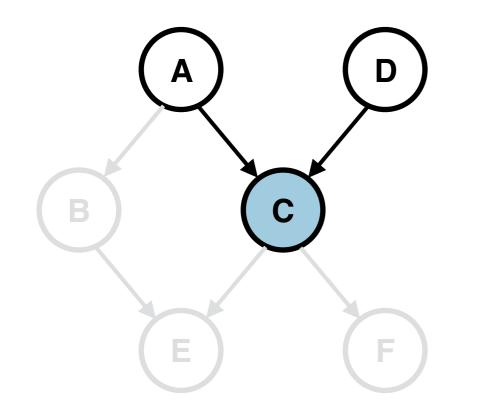
$$A \perp E \mid B$$

Case #2: Tail to Tail



$$\mathbb{P}(C, E, F) = \mathbb{P}(C)\mathbb{P}(E \mid C)\mathbb{P}(F \mid C) \Rightarrow$$
$$\mathbb{P}(E, F \mid C) = \frac{\mathbb{P}(C)\mathbb{P}(E \mid C)\mathbb{P}(F \mid C)}{\mathbb{P}(C)}$$
$$= \mathbb{P}(E \mid C)\mathbb{P}(F \mid C)$$
$$\checkmark$$
$$E \perp F \mid C$$

Case #3: Head to Head -----> Also known as colliders



$$\mathbb{P}(A, C, D) = \mathbb{P}(A)\mathbb{P}(D)\mathbb{P}(C \mid A, D) \Rightarrow$$

$$\mathbb{P}(A, D \mid C) = \frac{\mathbb{P}(A)\mathbb{P}(D)\mathbb{P}(C \mid A, D)}{\mathbb{P}(C)}$$

$$= \mathbb{P}(A, D \mid C)$$

$$\clubsuit$$

$$A \not\perp D \mid C$$

Explaining Away:

- A: Earthquake
- B: Break-in
- **C:** Motion alarm

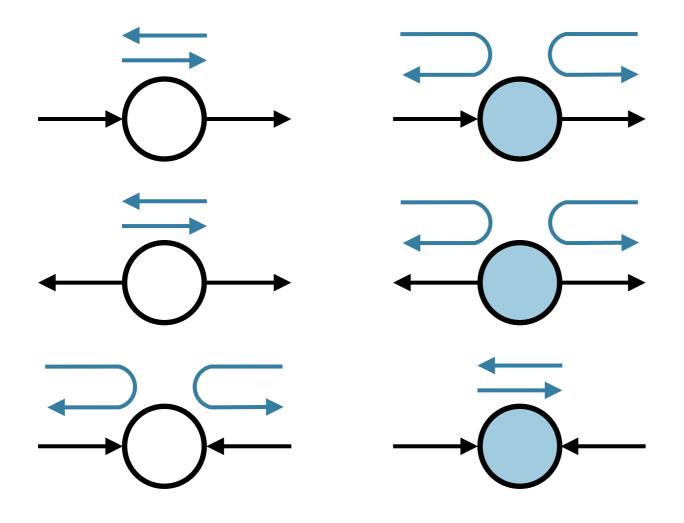
Given that the motion alarm went off, if we learn that a burglar broke in, then we know that it's unlikely an earthquake happened The burglary event "explains away" the earthquake event

D-Separation through **Bayes Ball**

This might be a little complicated to remember and apply, so let's look at an easier way to work out d-separation

D-Separation through **Bayes Ball**

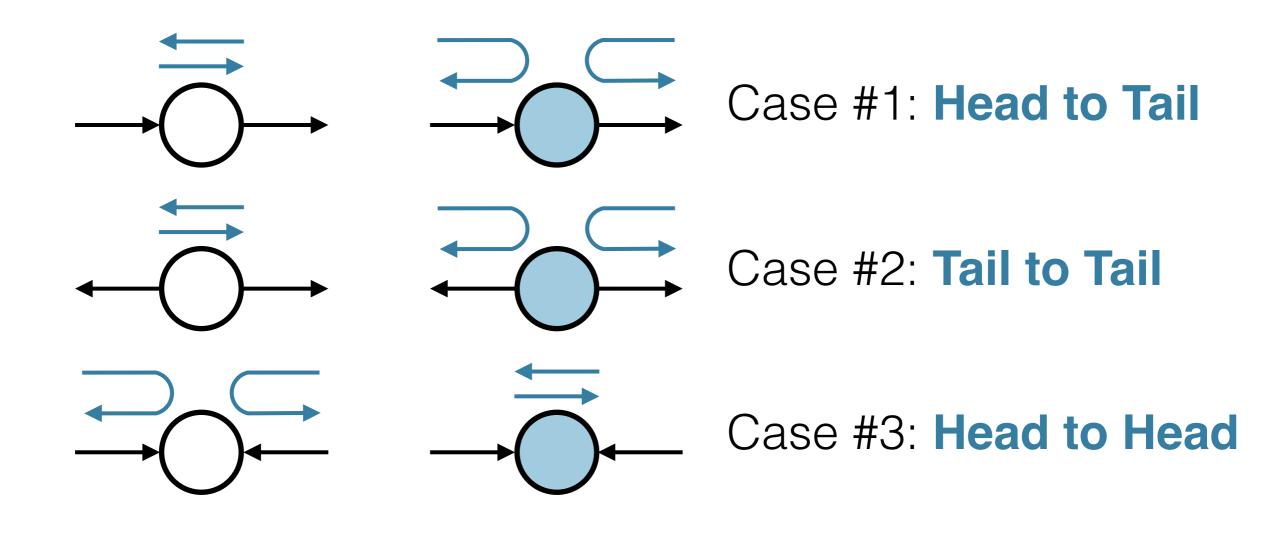
Imagine a ball, **Bayes ball**. This ball is allowed to "travel" on our model, but only in certain **allowed ways**:



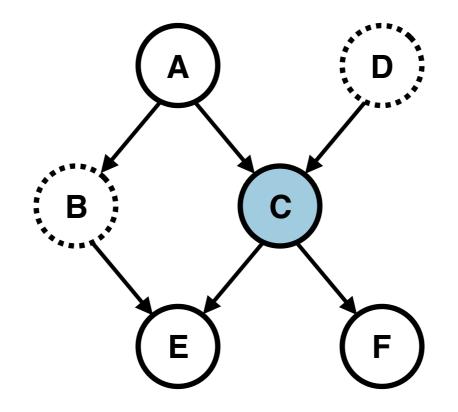
Two variables are conditionally independent when Bayes ball cannot travel from one to the other

D-Separation through **Bayes Ball**

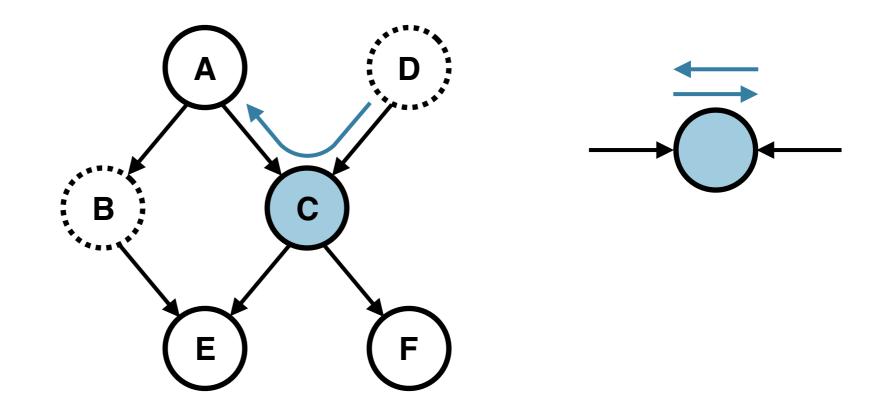
Imagine a ball, **Bayes ball**. This ball is allowed to "travel" on our model, but only in certain **allowed ways**:



Is **B** independent of **D** given **C**?

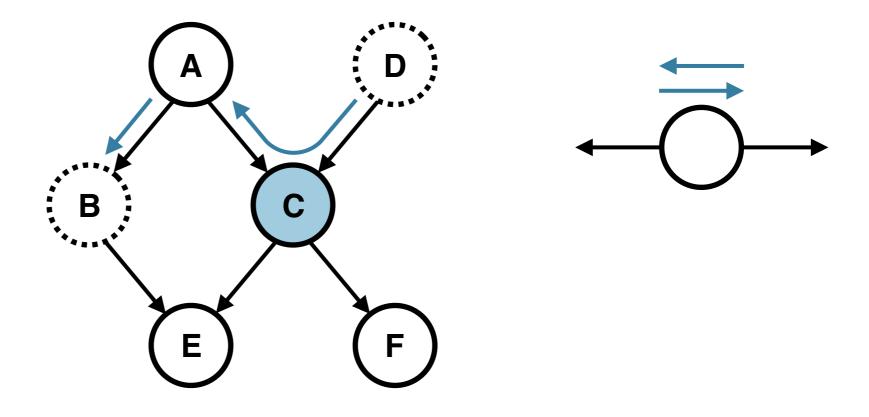


Is **B** independent of **D** given **C**?

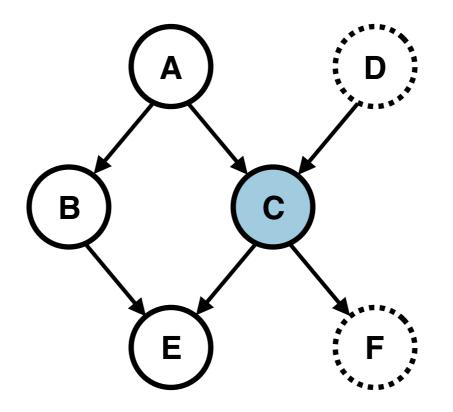


Is **B** independent of **D** given **C**?

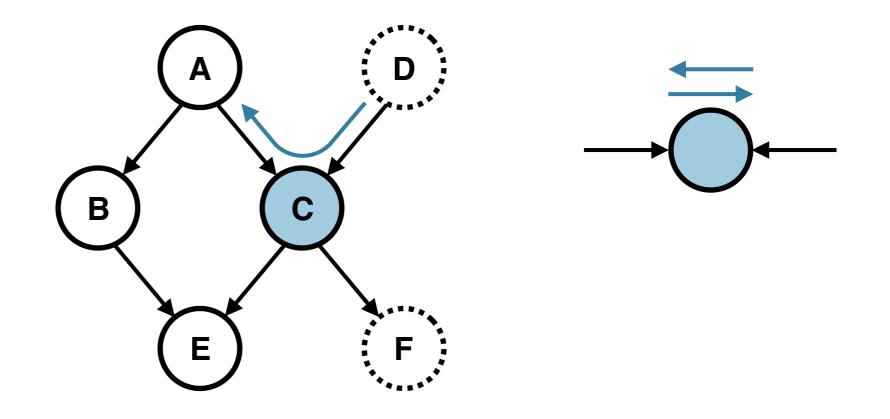
Bayes ball could reach **D** from **B** and thus the answer is **NO**



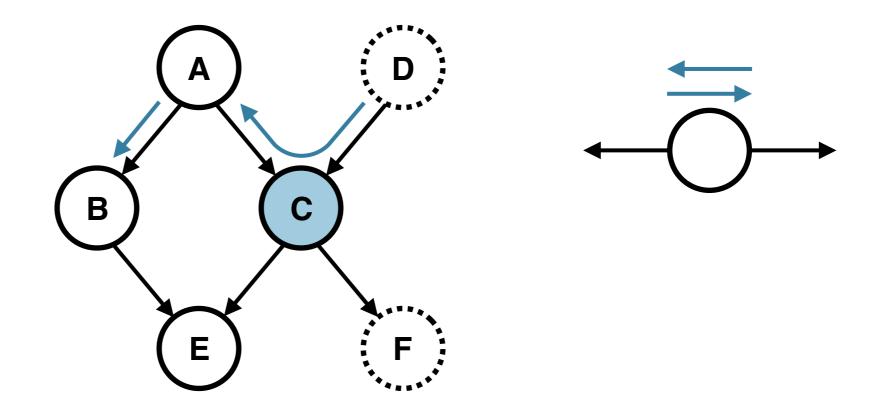
Is **F** independent of **D** given **C**?



Is **F** independent of **D** given **C**?

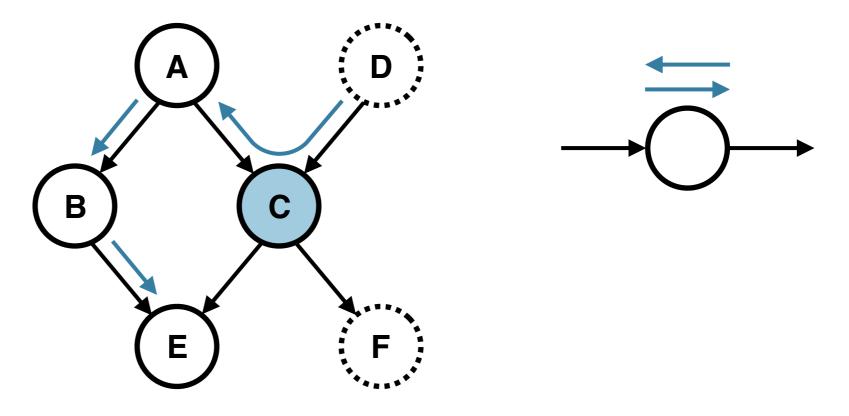


Is **F** independent of **D** given **C**?



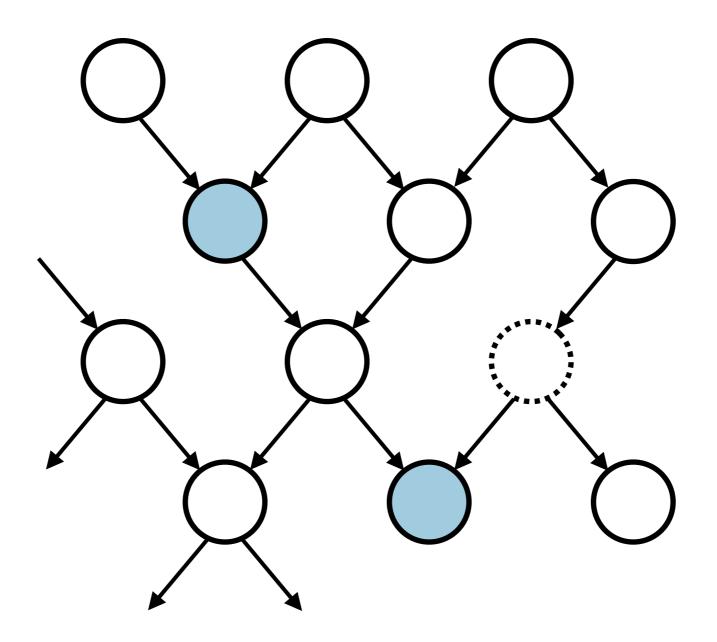
Is **F** independent of **D** given **C**?

Bayes ball got stuck going from **D** from **F** and thus the answer is **YES**



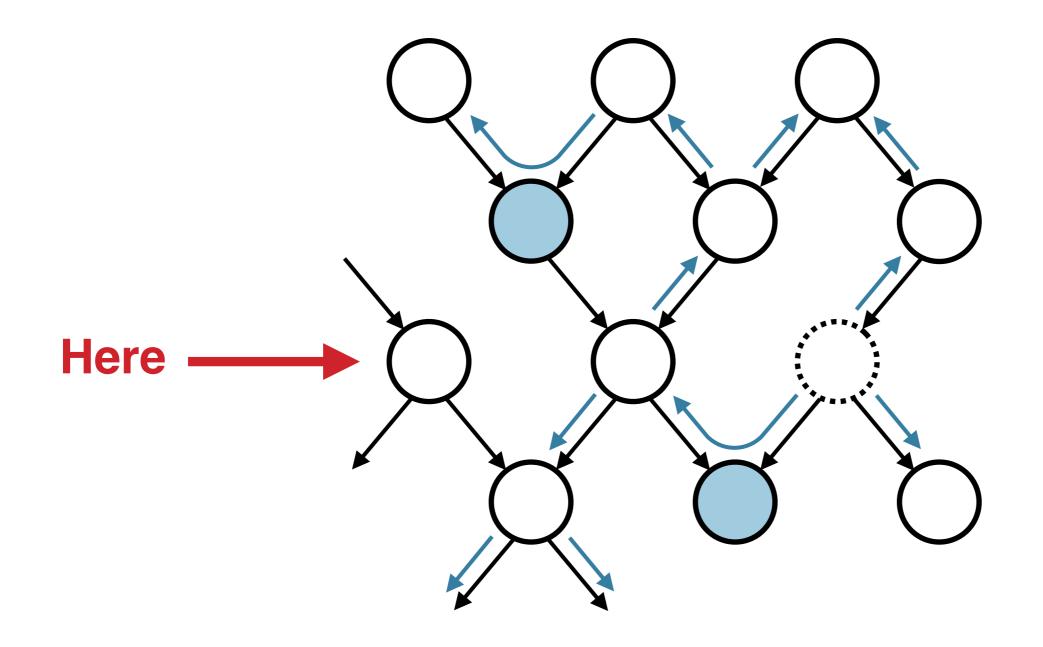
A More Complicated Example

Where can Bayes ball not go to, starting from the dashed variable?

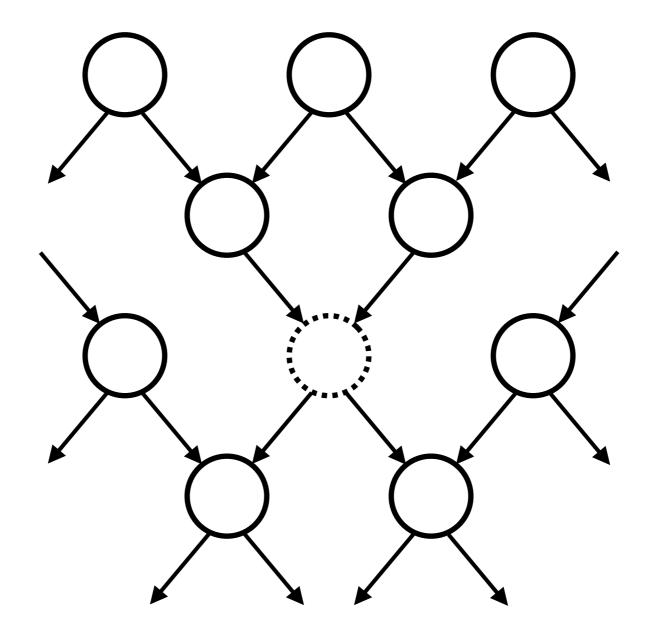


A More Complicated Example

Where can Bayes ball not go to, starting from the dashed variable?

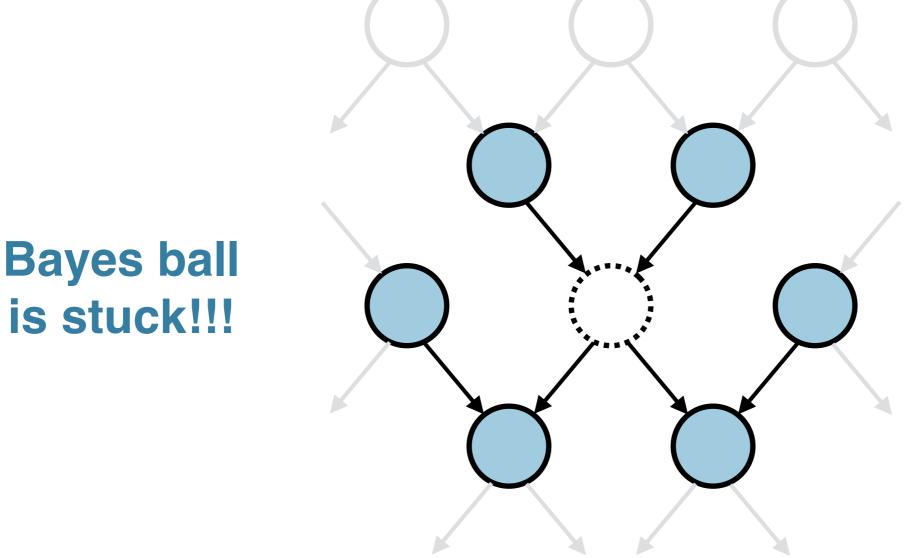


Markov Blanket



Markov Blanket

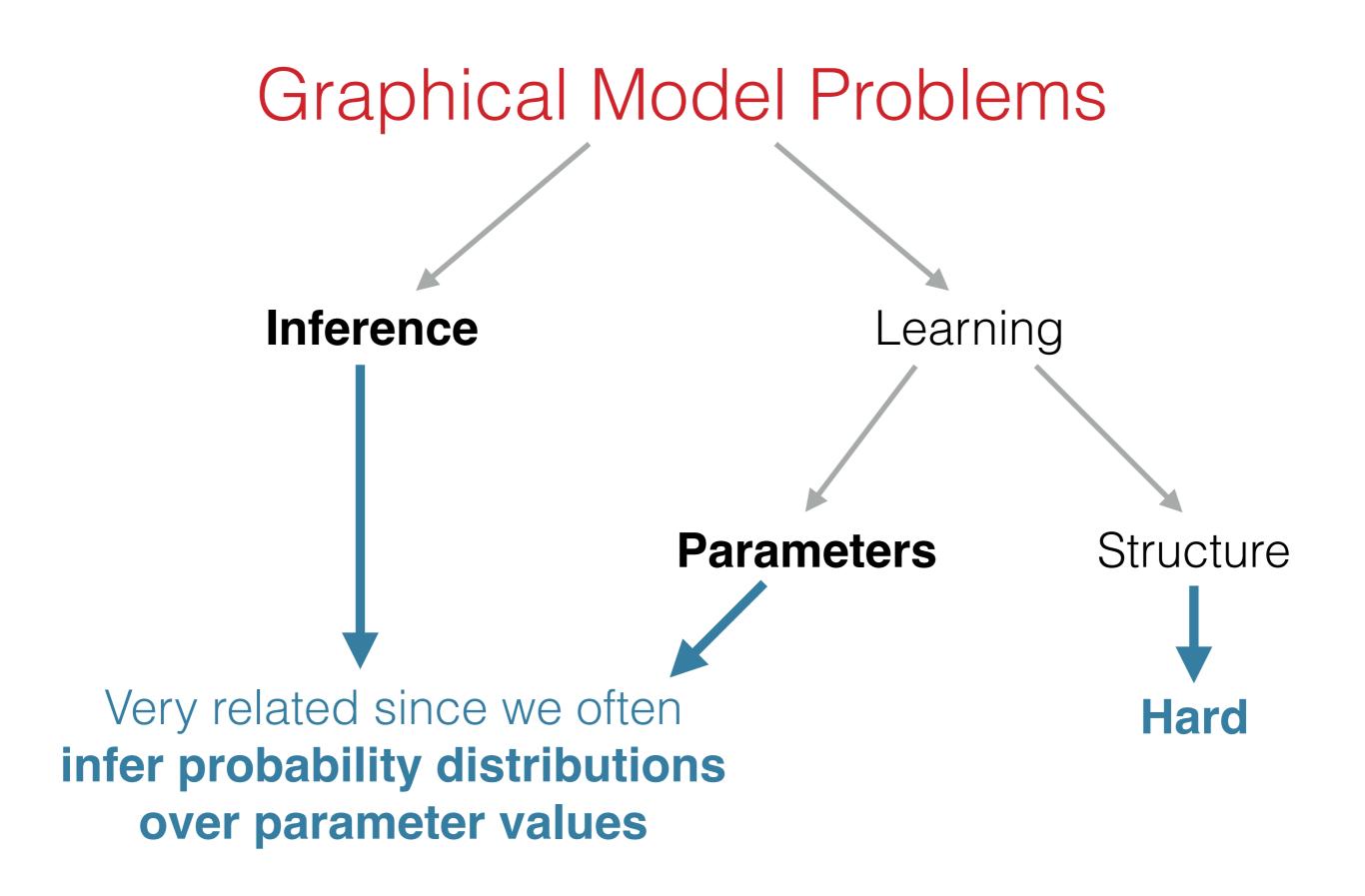
A variable is independent of every other variable in the model, given its **parents**, its **children**, and its **children's parents**



That is called the Markov blanket of a variable

Dependencies...so what?

We talked so much about understanding when there are dependencies/independencies between variables and how to model them, but why are they important?



Inference in Graphical Models

Probability of joint assignment over all **n** variables is **easy** — we just have to lookup the conditional probability tables (O(n) complexity)

Marginal probability distribution of a single variable is generally **hard** — we need to sum over all possible assignments to the rest of the variables ($\mathcal{O}(|\mathcal{D}|^n)$ complexity)

$$\mathbb{M}(X_i) = \sum_{X_1 \in \mathcal{D}_1} \cdots \sum_{X_{i-1} \in \mathcal{D}_{i-1}} \sum_{X_{i+1} \in \mathcal{D}_{i+1}} \cdots \sum_{X_n \in \mathcal{D}_n} \mathbb{P}(X_1, \dots, X_n)$$

Conditional independencies help us "move those sums around" and reduce complexity

Marginal Example

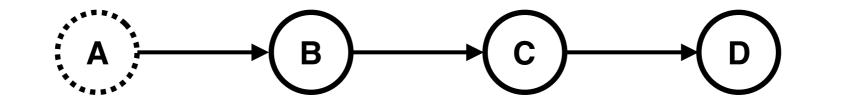
Linear Chain

С В D Α

 $\mathbb{P}(A, B, C, D) = \mathbb{P}(A)\mathbb{P}(B \mid A)\mathbb{P}(C \mid B)\mathbb{P}(D \mid C)$

Marginal Example

Linear Chain



 $\mathbb{P}(A, B, C, D) = \mathbb{P}(A)\mathbb{P}(B \mid A)\mathbb{P}(C \mid B)\mathbb{P}(D \mid C)$

 $\mathbb{M}(A) = \mathbb{P}(A) \sum_{B \in \operatorname{Val}(B)} \mathbb{P}(B \mid A) \sum_{C \in \operatorname{Val}(C)} \mathbb{P}(C \mid B) \sum_{D \in \operatorname{Val}(D)} \mathbb{P}(D \mid C)$

Each step costs $|\mathcal{D}|^2$ operations and so the complexity is now quadratic:

 $\mathcal{O}(n|\mathcal{D}|^2)$

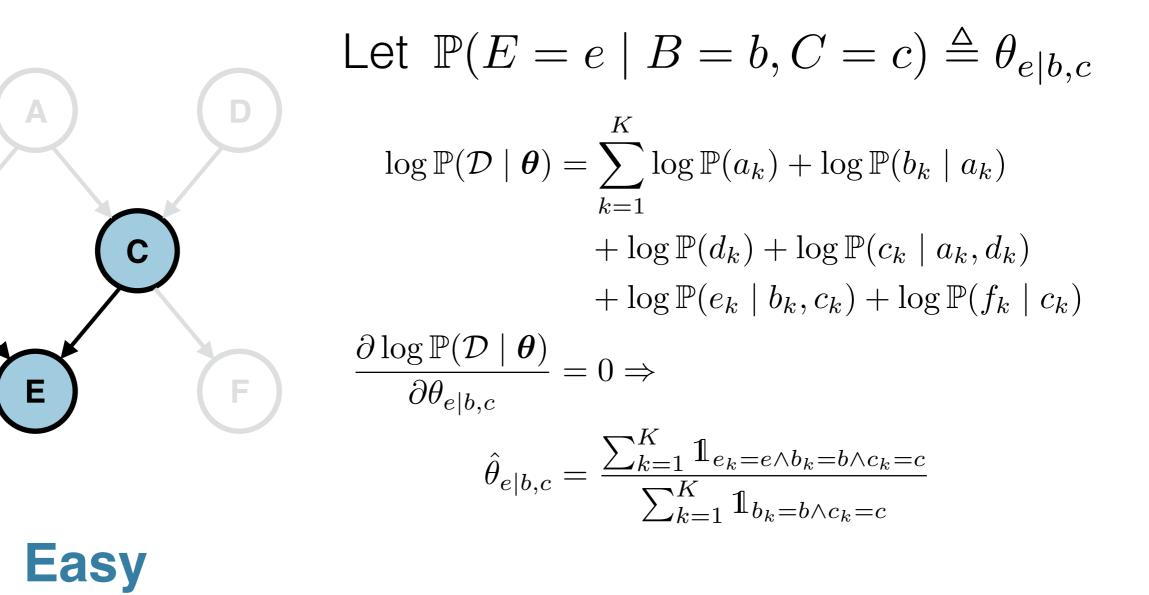
Learning in Graphical Models

	Fully Observed Variables	Partially Observed Variables
Known Structure	Easy	Interesting Frequent
Unknown Structure	Doable	Hard

Learning in Graphical Models

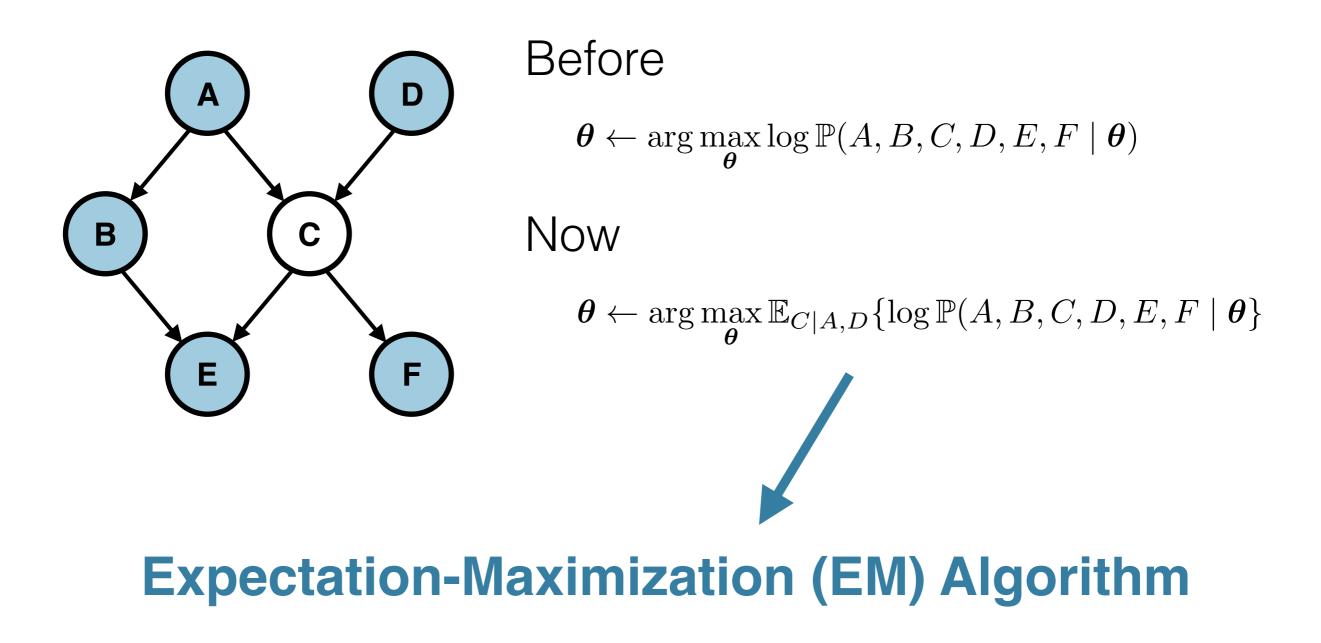
We want to estimate the model parameters for a **known** structure and with fully observed data

В

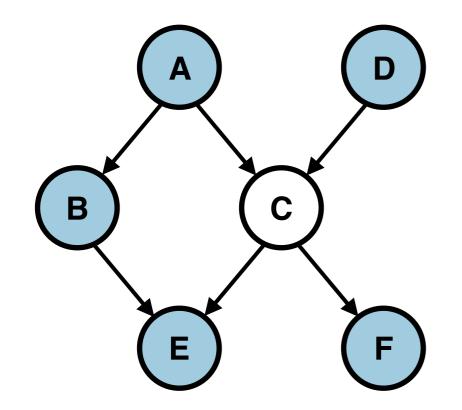


Learning in Graphical Models

We now want to estimate the model parameters for a **known structure** and with **partially observed data**



We begin with an arbitrary choice for our parameters and iterate over the following steps, until convergence



Guaranteed to find a local maximum

E Step: Estimate the values of the unobserved variables using the current parameters

M Step: Use the observed variables along with our estimates of the unobserved variables from the previous E step to compute a maximum likelihood estimate for our parameters and update them

We begin with an arbitrary choice for our parameters and iterate over the following steps, until convergence

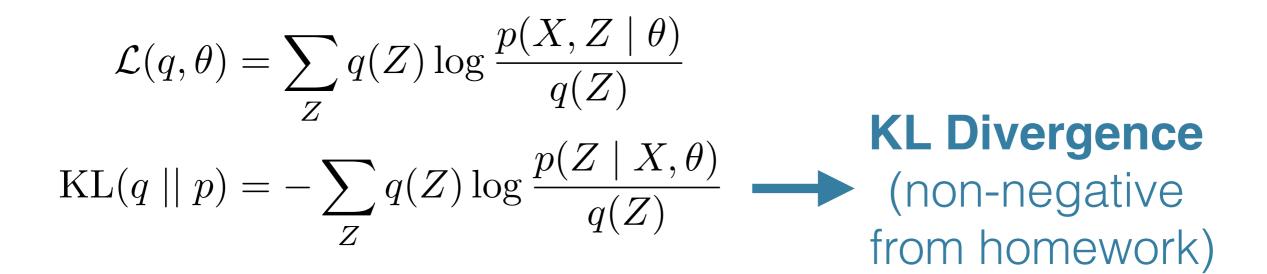
E Step D $\mathbb{E}\{C \mid A = a, D = d\} = \mathbb{P}(C = 1 \mid A = a, D = d)$ $= \theta_{C=1|a,d}$ B M Step F $\hat{\theta}_{c|a,d} = \frac{\sum_{k=1}^{K} \mathbb{1}_{c_k = 1 \land c_k = c \land d_k = d} \mathbb{E}\{C \mid A = a_k, D = d_k\}}{\sum_{k=1}^{K} \mathbb{1}_{c_k = c \land d_k = d}}$ Ε $=\frac{\sum_{k=1}^{K} \mathbb{1}_{c_k=1 \wedge c_k=c \wedge d_k=d} \theta_{C=1|a,d}}{\sum_{k=1}^{K} \mathbb{1}_{c_k=1 \wedge c_k=c \wedge d_k=d} \theta_{C=1|a,d}}$ **Guaranteed to find** a local maximum

In the simple **discrete variable** case, we do the same thing as in **maximum likelihood estimation**, but we simply **replace each unobserved variable count by its expected count**

EM is not exact, but why is it guaranteed to find a local maximum?

Intuition: The following holds for any distribution q(Z)

 $\log p(X \mid \theta) = \mathcal{L}(q, \theta) + \mathrm{KL}(q \mid \mid p)$

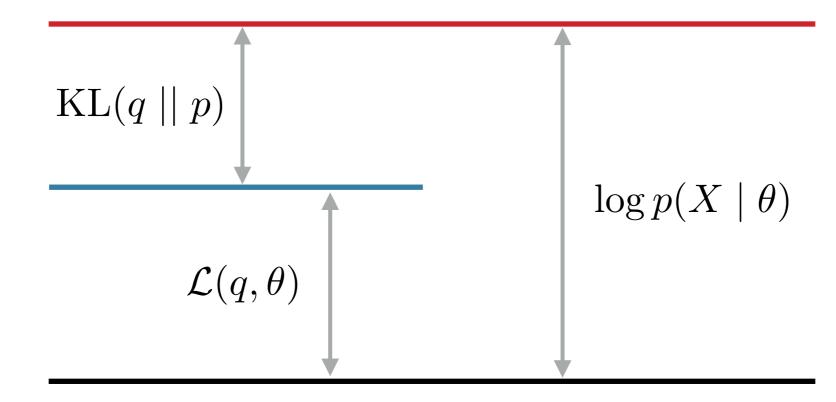


X: observed variables Z: unobserved variables

Intuition: The following holds for any distribution q(Z)

$$\log p(X \mid \theta) = \mathcal{L}(q, \theta) + \mathrm{KL}(q \mid \mid p)$$

$$\mathcal{L}(q,\theta) = \sum_{Z} q(Z) \log \frac{p(X, Z \mid \theta)}{q(Z)} \le \log p(X \mid \theta)$$

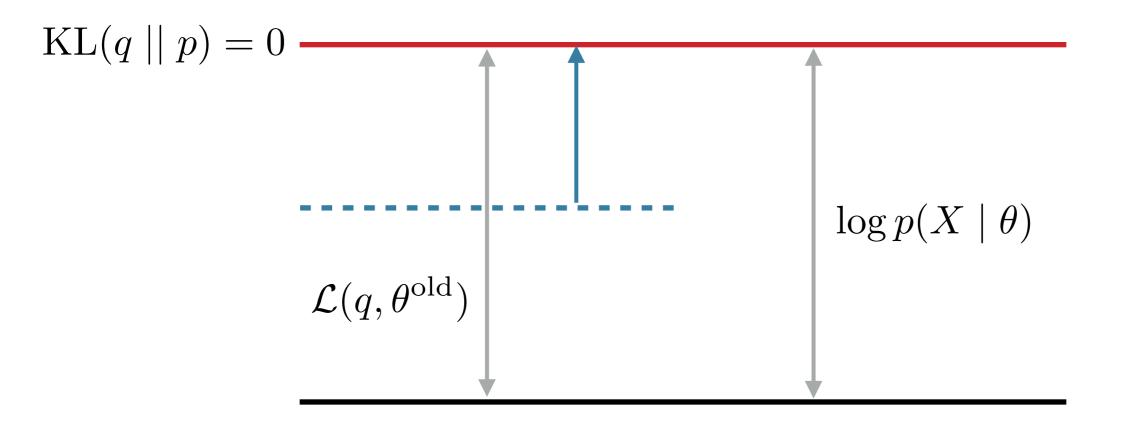


Visualization idea borrowed from Chris Bishop's book

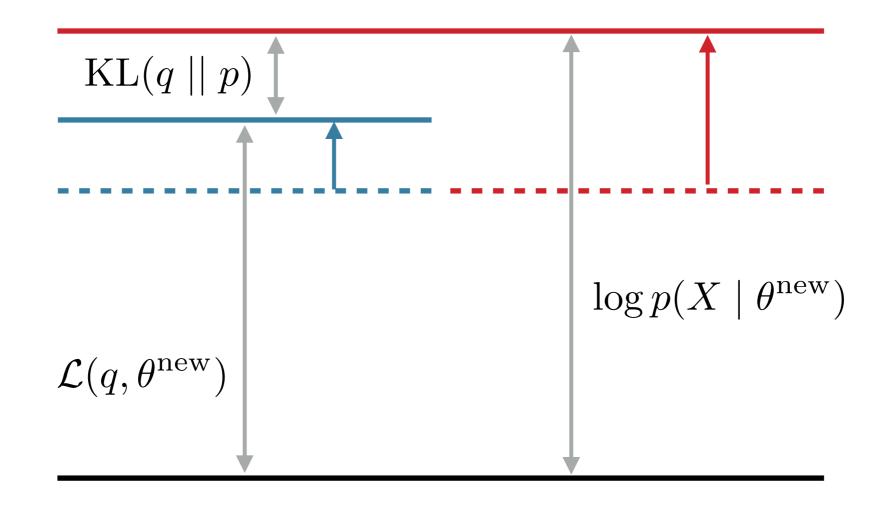
E Step Intuition: We maximize $\mathcal{L}(q, \theta^{\text{old}})$ while holding θ^{old} fixed. We thus set

$$q(Z) = p(Z \mid X, \theta^{\text{old}})$$

Remember that $\log p(X \mid \theta) = \mathcal{L}(q, \theta) + \mathrm{KL}(q \mid \mid p)$



M Step Intuition: We maximize the lower bound with respect to θ while holding q(Z) fixed



Markov Chain Monte Carlo (MCMC)

Conditional probability distributions when dealing with a known structure and partially observed variables can be computed using sampling — Markov Chain Monte Carlo (MCMC) methods are often used

Gibbs sampling is a very common such method

- Initialize unobserved variables to some values
- Sample each unobserved variable in sequence, while fixing the rest to their last sampled value
- Burn (i.e., throw away) the first few samples and then thin the rest (e.g., keep every 10th sample)

The distribution of the samples converges to the true posterior distribution of the unobserved variables

Bayesian Network Structure Learning

Learning structure is not that easy

- In general requires lots of data (can **overfit** easily)
- Huge search space we use priors to constrain it

But there exist some algorithms for certain special cases

e.g., Chow-Liu for tree structures

Next time

Finds the tree structure that minimizes KL divergence (i.e., mutual information)

Questions?